# Statistical Catch-at-Length Assessment of S. mentella and S. fasciatus in Units 1+2 

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April 2014


#### Abstract

Summary Assessments of the two redfish species in these Units is attempted simultaneously to take account of the fact that they are distinguished only in the survey results and not in the commercial catches. Fitting the declines in the survey indices in Unit 1 for the earlier years proves a particular problem, and leads to the question of whether bounds could be placed on survey catchabilities $(q)$ to avoid what seem to be some unrealistically high estimates of $q$ for $S$. fasciatus. Allowing for occasional large recruitments in these populations shows promise for improving the fits to those survey indices. However this needs further investigation to determine whether associated poor fits to the survey catch-at-length data can be avoided, and whether estimation stability can be improved.


## Introduction

This document presents results from an application of a Statistical Catch-at-Length (SCAL) assessment approach to the S. mentella and S. fasciatus resources in Units 1+2. Because (unlike the survey data) the commercial catch data available for this region is speciesaggregated, the approach assesses both species simultaneously so as to be able to fit to these species-combined data.

The results presented in this document fall into two sections. First there are those for some initial runs which were discussed at a teleconference in early March 2014 (though slightly modified for reasons given below). Following that teleconference, ideas for further runs were offered and subsequently developed, and those follow in a second section.

## Data and Methods

The data used are listed in Appendix A.
The methodology, detailed in Appendix $B$, is broadly as described in Rademeyer and Butterworth (2011), with some key features and changes described below:

1) An age-structured model is used rather than an age-aggregated approach (as in McAllister and Duplisea, 2012) for a more realistic representation of the dynamics.
2) The new Campana ageing data are used: for each species, a von Bertalanffy growth curve through the origin has been fitted to these data and the resulting parameters used in the assessment.
3) Instead of assuming a knife-edged maturity-at-age 9, a knife-edged maturity-at-length 22 cm is assumed which is then converted to maturity-at-age using the estimated length-at age distributions.
4) Although the survey biomass index is taken as the mature biomass only ( $\geq 22 \mathrm{~cm}$ ), the model is now fitted to the whole range of survey catch-at-length data available.
5) Because the commercial catches and catch-at-length data are not disaggregated by species, the assessment models both species simultaneously.
6) No assumption about the species split of the catches is made on input; rather flexibility is allowed in the model by estimating the annual S. fasciatus proportion in the catches directly, by means of the following penalty added to the negative log-likelihood:
$-\ell n L^{F P p e n}=\sum_{u} \sum_{y=1960}^{2009}\left[\frac{\left(p_{y}^{u}-\mu^{u}\right)^{2}}{2\left(\sigma^{u}\right)^{2}}\right]$
where
$p_{y}^{u}$ is the estimated proportion of $S$. fasciatus in the catch in year $y$ and Unit $u$,
$\mu^{u}$ and $\sigma^{u}$ are the mean and standard deviation respectively of the distribution of $S$. fasciatus proportions in Unit $u$ based on the survey species split information (McAllister and Duplisea, 2012). For Unit 1, $\mu^{u}=0.40$ and $\sigma^{u}=0.16$, and for Unit 2, $\mu^{u}=0.53$ and $\sigma^{u}=0.10$.
7) Scenarios with occasional high recruitments are implemented by allowing a large variability about the stock-recruitment relationship ( $\sigma_{R}=1.5$ ), essentially permitting the recruitments to be estimated freely.
8) For the scenarios with a change in carrying capacity, the changes are modelled as a random walk (separately for each species):
$K_{y}=K_{y-1} 1^{\varepsilon_{y}}$
with the following penalty added to the negative log-likelihood:
$-\ell n L^{\text {Kpen }}=\sum_{y=1961}^{2009}\left[\varepsilon_{y}^{2} / 2 \sigma_{K}^{2}\right]$
with $\varepsilon_{y}$ estimated in the model fitting procedure and $\sigma_{K}=0.3$.
9) A penalty on the survey catchability coefficients is used for all scenarios in the spirit of a prior to avoid the results going into implausible regions of parameter space (particularly S. fasciatus' survey catchability $q$ going unrealistically high). The following penalty is added to the negative log-likelihood to effect this:
$-\ell n L^{q p e n}=\sum_{i=1}^{n_{\text {surves }}}\left[\frac{2\left(q^{i}-l_{b}\right)}{\left(u_{b}-l_{b}\right)}-1\right]^{p}$
with the upper and lower bounds ( $l_{b}$ and $u_{b}$ ) chosen as: $l_{b}=0.2, u_{b}=2.0$ and $p=16$
10) The catch-at- length data are downweighted by a factor of 0.01 instead of 0.1 in Rademeyer and Butterworth (2011). This is to ensure that catch-at-length information does not unduly influence the model's attempt to fit the survey index data.
11) In the cases where log-normally distributed fluctuations about the stock-recruitment relationship are admitted, and with a high value for the extent of variability $\sigma_{R}=1.5$ to allow for the possibility of occasional very large recruitments, the starting abundance and age-structure corresponds to median rather than to mean recruitment (and carrying capacity $K$ similarly) so that this reflects the typical situation absent those large year classes.

## Results

First results for four runs are presented:

1) " $\sigma_{R}=0.4$ ": does not allow for occasional high recruitment or changes in carrying capacity.
2) " $\sigma_{R}=1.5$ ": allows for occasional high recruitment.
3) " $\sigma_{R}=0.4$, with changes in $K^{\prime}$ : makes allowance for changes in the carrying capacities for each species.
4) " $\sigma_{R}=0.4$, with changes in $K$, more weight 1990-1995 surveys": as 3 ) above, but with more weight ( $W=10$ ) added to the 1990-1995 survey data points. This scenario was selected because no other run fits the early survey index declines.

Table 1 compares the results for these four scenarios. Note that compared to material circulated for the March teleconference, results for 2 ) differ because of the modification indicated in note 11) above - consequently this is termed scenario $2 a$ in this paper. Furthermore results for 4) also differ slightly because of an earlier error in the value accorded to $W_{\text {CAL }}$ for this run.

The fits to the survey biomass indices for S. mentella and S. fasciatus are plotted in Figure 1 for each of the four scenarios, while Figure 2 compares the spawning biomass and recruitment trajectories.

Figures 3 to 5 give more detailed results for scenario 2 a ( $\sigma_{R}=1.5$ ). Figure 3 plots the catch trajectories by species and Unit as well as the estimated S. fasciatus proportion in the catch for this scenario. The estimated survey and commercial selectivities are shown in Figure 4. The commercial selectivities are taken to be the same for Unit 1 and Unit 2. Finally Figure 5 plots the fits to the survey biomass and catch-at-length data.

At the March 2014 teleconference, a further series of scenarios were suggested and some more have been added by the authors. The corresponding runs have been based on the $\sigma_{R}=1.5$ run (scenario $2 a-$ as this was considered at the teleconference to hold the most promise) for all except one scenario (scenario 5). Scenario 5 is based on run1 ( $\sigma_{R}=0.4$ ) because that yields values for $q$ which are closer to the Bundy estimate than are the $q$ values for run 2a ( $\sigma_{R}=1.5$ ).
5) Fixed $q=0.43$ (as advised to correspond to the estimate by Alida Bundy).
6) Lessen the prior constraints on $q$ (bounds changed to 0.1 to 5).
7) Flat survey selectivity from length 30 cm onwards.
8) Alternative priors for the species split of the catches, keeping the standard deviation as in run 2 a. for Unit 1, $\mu^{1}=0.60$ and for Unit 2, $\mu^{2}=0.73$ and b. for Unit $1, \mu^{1}=0.20$ and for Unit 2, $\mu^{2}=0.33$.
9) Logistic survey selectivities.
10) Allow for large recruitment variability, forcing a fit to the early survey declines.

Tables 2 and 3 compare results for the scenarios described above. Results for scenario 5 (fixed $q=0.43$ ) are shown in Table 2 together with the corresponding scenario 1 results, while results for the other scenarios are given in Table 3.

The fits to the survey biomass indices for S. mentella and S. fasciatus are plotted in Figures 6 and 7 for the scenarios described above.

The fits to the survey CAL data for scenarios $2 \mathrm{a}\left(\sigma_{R}=1.5, W_{C A L}=0.01\right.$ ) and 2 b ( $\sigma_{R}=1.5$, $W_{C A L}=0.1$ ) are compared in Figure 8.

Figure 9 plots the commercial and survey selectivities-at-length for scenario 7, for which the survey selectivities are forced flat from length 30 cm onward.

Estimated species-disaggregated catch trajectories and S. fasciatus proportions in the catches are compared in Figure 10 for scenarios $2 a, 8 a$ and $8 b$, which assume different priors for the species split of the catches.

## Discussion

First the results reported for the teleconference and the impact of some subsequent adjustments to those are discussed.

1) Previously it was reported that occasional high recruitment or changes in $K$ could partly explain the earlier S. fasciatus survey results, but not the S. mentella ones. Hence run 4) is introduced, "forcing" the model to fit both those Unit 1 early declines. With the modification of methodology note 11) to commence with , the occasional high recruitment option of run 2a) can explain the earlier S. mentella decline through a large increase in the estimated size of the 1981 year class, but there are some associated problems as discussed further below.
2) All of runs 2a) to 4) set high values for survey catchability for S. fasciatus in Unit 2 in particular in an attempt to reduce estimates of recent biomass so as to be able to better reflect the decline in the S. fasciatus survey index in Unit 1 over 1990-1995.
3) Run 2 a ), with high $\sigma_{R}=1.5$ to allow for occasional high recruitment pulses (as indeed are then estimated - see Fig. 2), leads to a higher estimate of $K$ for S. mentella which is less depleted relative to $K$, though $S$. fasciatus is more depleted relative to $K$. (this again differs from the results reported for the teleconference as a consequence of method modification 11).
4) Forcing the varying $K$ scenario to fit the early survey index declines in Unit 1 in run 4) leads to higher estimated initial $K$ and some higher survey catchability $q$ values for both species, The current biomass estimated for S. mentella is appreciably less.

Specifically in relation to the Figures shown for scenario 2 only (though these points also apply to the other scenarios):

- There seems little information in the process to update the "prior" on the S. fasciatus proportion of the catch appreciably (Figure 3).
- The low survey selectivities in the $20-30 \mathrm{~cm}$ range (Figure 4 ) are surprising. The follow from the observed length frequency distributions. What mechanism is responsible for the absence of these lengths in the survey data?

The following are features of interest in the results for the sensitivities suggested at the teleconference and related further runs.
5) Scenario 2a) seems very promising in showing an ability to fit the initial decline in the survey index for $S$. mentella for Unit 1 by estimating a very large 1981 year-class (Figures 1 and 2). However, that is at the cost of a severe misfit to the corresponding CAL data (Figure 5), with an absence in the model of the larger S. mentella observed in these surveys. If the weight on these CAL data is increased (scenario $2 b$ ), they are fitted much better (Figure 8), but then the initial decline in the index is no longer reflected (Figure 6). Thus basically there is a conflict between these two data sources given the current model, which further work should attempt to resolve.
6) Setting $q=0.43$ (scenario 5) makes effectively no difference to the results from the comparative scenario 1 with $\sigma_{R}=0.4$ (Table 2 and compare Figures 6 and 1).
7) Scenario 6 which widens the constraints on the range for $q$ leads to a slightly better fit to the early decline in the abundance index for $S$. fasciatus in Unit 1 (Figure 6), but the estimates of the corresponding survey $q$ 's become extremely high (Table 3).
8) Flattening survey selectivity above a length of 30 cm (scenario 7) makes little difference to estimates of importance for management (Table 3).
9) Changing the prior for the species split of the catch to reflect a bigger proportion of $S$. fasciatus (scenario 8a) improves the fit overall (slightly) and particularly that to the early Unit 1 S. fasciatus survey index decline (Table 3 and Figure 7). A change in the prior in the other direction leads to a deterioration in the fit to this early decline for both species.
10) A logistic form for the survey selectivities (scenario 9) leads to an appreciable deterioration in the fits to the S. mentella survey index of abundance (Table 3 and Figure 7).
11) Forcing a fit to the early survey indices in Unit 1 for both species (scenario 10, Figure 7) leads to a deterioration in the fit to the commercial CAL data in particular, but does suggest that the $S$. fasciatus population is less depleted relative to $K$ (Table 3).

Overall, though the possibility of allowing for occasional high recruitments shows promise, in particular as a means of accounting for the initial declines in the survey indices in Unit 1 for both species, we must stress that convergence of the model fit is difficult to achieve for this approach. More work is needed to improve the estimation stability of this approach before the results which it provides could be regarded as reliable.

## Further work

Issues meriting further discussion include:

- Restrictions (particularly upper bounds) to be placed on the survey catchabilities $q$.
- The importance of an assessment reflecting the 1990-1995 declines in the survey index for both species in Unit 1.
- Possible further spatial sub-structuring of the assessment to be able to accommodate (inter alia) further surveys with only partial coverage of a Unit.
- Alternative approaches to modelling occasional high recruitments (mixture distributions perhaps?), and whether there is further information that might be included to assist stablisethe assessment when such possibilities are admitted.


## References

McAllister M and Duplisea DE. 2012. Production model fitting and projection for Acadian redfish (Sebastes fasciatus) in Units 1 and 2. DFO Can. Sci. Advis. Sec. Res. Doc. 2012/103. iii + 34p
Rademeyer RA and Butterworth DS. 2011. Initial applications of statistical catch-at-age assessment methodology to Atlantic redfish. Document submitted to Canadian ZAP meeting related to Precautionary Approach reference points for redfish populations, Mont-Joli, October 2011: 34pp.

Table 1: Results of fits of scenarios 1 to 4 for redfish in Units $1+2$. Values fixed on input rather than estimated are shown in bold. Mass units are ' 000 t . In cases where the value of the pre-exploitation spawning biomass $K$ changes within the assessment period, the two columns for $K^{s p}$ reports the carrying capacity in the first and last year of the assessment period respectively. The value of $W_{\text {CAL }}$ is 0.01 for all these runs.

|  | 1) $\sigma_{R}=0.4$ |  | 2a*) new $\sigma_{R}=1.5$ |  | 3) $\sigma_{R}=0.4$, with changes in $K$ |  | 4) $\sigma_{R}=0.4$, with $K$ changes, more weight 1990-1995 surveys |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| -InL: overall | -157.9 |  | -53.3 |  | -180.8 |  | -247.7 |  |  |  |
| -InL: survey | 13.0 |  | -18.6 |  | -13.0 |  | -108.6 |  |  |  |
| -InL: survCAL | -38.8 |  | -35.0 |  | -39.4 |  | -33.7 |  |  |  |
| -InL: comCAL | -44.3 |  | -47.7 |  | -43.1 |  | -32.2 |  |  |  |
| -InL: catchpen | 0.00 |  | 0.00 |  | 0.00 |  | 0.00 |  |  |  |
| -InL: FascProppen | 0.18 |  | 0.22 |  | 0.25 |  | 5.78 |  |  |  |
| -InL: SRpen | -88.53 |  | 47.56 |  | -89.58 |  | -88.88 |  |  |  |
| -InL: qpen | 0.56 |  | 0.26 |  | 0.13 |  | 1.45 |  |  |  |
| -InL: Kpen |  |  | 3.86 | 8.42 |  |  |  |
|  | S. mentella S.fasciatus |  |  |  | S. mentella | S. fasciatus | S. mentella S.fasciatus |  | S. mentella S. fasciatus |  |  |  |
| $h$ | 0.67 | 0.67 | 0.67 | 0.67 | 0.67 | 0.67 | 0.67 |  | 0.67 |  |
| M | 0.100 | 0.125 | 0.100 | 0.125 | 0.100 | 0.125 | 0.100 |  | 0.125 |  |
| $\theta$ | 1.00 | 1.00 | 1.00 | 1.00 | 1.00 | 1.00 | 1.00 |  | 1.00 |  |
| $\zeta$ | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 |  | 0.00 |  |
| $K^{5 p}$ | 1299 | 724 | 1741 | 704 | 109072 | 1156281 | 2039 | 202 | 2228 | 539 |
| $B^{\text {Sp }}{ }_{2009}$ | 946 | 519 | 1599 | 132 | 101 | 124 | 60 |  | 174 |  |
| $B^{5 P}{ }_{2009} / K^{5 P}$ | 0.73 | 0.72 | 0.92 | 0.19 | 0.091 .40 | $0.11 \quad 0.44$ | 0.03 | 0.30 | 0.08 | 0.32 |
| MSYL ${ }^{\text {sp }}$ | 0.3 | 0.4 | 0.30 | 0.30 | 0.31 | 0.30 | 0.32 |  | 0.31 |  |
| $B^{\text {Sp }}{ }_{M S Y}$ | 452.7 | 280.1 | 527.2 | 210.7 | 22.2 | 83.2 | 65.7 |  | 168.2 |  |
| MSY | 56.2 | 37.7 | 203.0 | 98.0 | 3.0 | 14.2 | 8.3 |  | 27.4 |  |
| Survey | $q$ 's $\sigma_{\text {Add }}$ | $q$ 's $\sigma_{\text {Add }}$ | $q$ 's $\sigma_{\text {Add }}$ | $q$ 's $\sigma_{\text {Add }}$ | $q$ 's $\sigma_{\text {Add }}$ | $q$ 's $\sigma_{\text {Add }}$ | $q$ 's | $\sigma_{\text {Add }}$ | $q$ 's | $\sigma_{\text {Add }}$ |
| Unit 1 | 0.250 .87 | 0.290 .99 | 1.490 .27 | $1.06 \quad 0.56$ | 0.620 .45 | 0.890 .54 | 1.49 | 0.16 | 1.93 | 0.30 |
| Unit 2 | $0.58 \quad 0.31$ | $0.68 \quad 0.32$ | 0.300 .13 | 1.880 .27 | 1.69 | 1.890 .30 | 1.98 | 0.00 | 1.96 | 0.27 |
| $\sigma_{\text {R_out }}$ | 0.07 | 0.06 | 0.69 | 0.49 | 0.03 | 0.02 | 0.05 |  | 0.06 |  |

* This is not identical to the run considered during the March teleconference, as the starting biomass corresponds to median rather than mean recruitment - see note 11) under Data and Methods.

Table 2: Results of fits of scenarios 1 and 5 (all q's fixed at 0.43 ) for redfish in Units $1+2$. Values fixed on input rather than estimated are shown in bold. Mass units are '000t.

|  | 1) $\sigma_{R}=0.4$ |  | 5) $\sigma_{R}=0.4, q=0.43$ |  |
| :---: | :---: | :---: | :---: | :---: |
| -InL: overall | -157.9 |  | -150.2 |  |
| -InL: survey | 13.0 |  | 16.8 |  |
| -InL: survCAL | -38.8 |  | -36.2 |  |
| -InL: comCAL | -44.3 |  | -42.6 |  |
| -InL: catchpen | 0.00 |  | 0.00 |  |
| -InL: FascProppen | 0.18 |  | 0.59 |  |
| -InL: SRpen | -88.53 |  | -88.88 |  |
| -InL: qpen | 0.56 |  | 0.13 |  |
| -InL: Kpen |  |  |  |  |
|  | S. mentella | S. fasciatus | S. mentella | S. fasciatus |
| $h$ | 0.67 | 0.67 | 0.67 | 0.67 |
| $M$ | 0.100 | 0.125 | 0.100 | 0.125 |
| $\theta$ | 1.00 | 1.00 | 1.00 | 1.00 |
| $\zeta$ | 0.00 | 0.00 | 0.00 | 0.00 |
| $K^{\text {sp }}$ | 1299 | 724 | 991 | 743 |
| $B^{\text {SP }} 2009$ | 946 | 519 | 688 | 536 |
| $B^{\text {Sp }}{ }_{2009} / K^{\text {Sp }}$ | 0.73 | 0.72 | 0.69 | 0.72 |
| MSYL ${ }^{\text {sp }}$ | 0.3 | 0.4 | 0.32 | 0.31 |
| $B^{5 p}{ }_{M S Y}$ | 452.7 | 280.1 | 317.1 | 230.7 |
| MSY | 56.2 | 37.7 | 40.2 | 37.1 |
| Survey | $q$ 's $\sigma_{\text {Add }}$ | $q$ 's $\sigma_{\text {Add }}$ | $q$ 's $\quad \sigma_{\text {Add }}$ | $q$ 's $\sigma_{\text {Add }}$ |
| Unit 1 | $0.25 \quad 0.87$ | 0.290 .99 | $0.43 \quad 0.99$ | 0.431 .00 |
| Unit 2 | $0.58 \quad 0.31$ | $0.68 \quad 0.32$ | $0.43 \quad 0.33$ | $0.43 \quad 0.31$ |
| $\sigma_{\text {R_out }}$ | 0.07 | 0.06 | 0.03 | 0.07 |

Table 3: Results of fits of scenarios $\mathbf{2 a}$ and $\mathbf{2 b}$, and $\mathbf{6}$ to $\mathbf{1 0}$ for redfish in Units $1+2$. Values fixed on input rather than estimated are shown in bold. Mass Units are ' 000 t. The value of $W_{\text {CAL }}$ is 0.01 unless otherwise indicated.



Figure 1: Fits to the survey biomass indices for scenarios 1 to 4


Figure 2: Spawning biomass and recruitment trajectories for $S$. mentella (blue lines) and $S$. fasciatus (red lines) for scenarios 1 to 4.


Figure 3: Catch trajectories by Unit and species, and estimated S. fasciatus proportion in the catch for scenario 2a ( $\sigma_{R}=1.5, \boldsymbol{W}_{C A L}=0.01$ ).


Figure 4: Commercial and survey selectivities-at-length for scenario $2 \mathrm{a}\left(\sigma_{R}=1.5, \boldsymbol{W}_{C A L}=0.01\right.$ ).


Figure 5: For scenario $\mathbf{2 a}\left(\sigma_{R}=1.5, \boldsymbol{W}_{C A L}=\mathbf{0 . 0 1}\right)$, fits to the survey biomass indices (first row), corresponding residuals (second row), fits to CAL data (as averaged over all the years for which data are available) (third row) and bubble plots of the standardised residuals for the fit to the CAL data (last row). The area of the bubble is proportional to the magnitude of the corresponding standardised residuals. For positive residuals the bubbles are blue/pink, whereas for negative residuals the bubbles are white.


Figure 6: Fits to the survey biomass indices for scenarios $\mathbf{2 a}$ and $\mathbf{2 b}$ and 5 to 7.


Figure 7: Fits to the survey biomass indices for scenarios 8a to 10.


Figure 8: For scenarios $\mathbf{2 a}\left(\sigma_{R}=1.5, \boldsymbol{W}_{C A L}=\mathbf{0 . 0 1}\right)$ and $\mathbf{2 b}\left(\sigma_{R}=1.5, \boldsymbol{W}_{C A L}=\mathbf{0 . 1}\right)$, fits to CAL data (as averaged over all the years for which data are available) (third row) and bubble plots of the standardised residuals for the fit to the CAL data (last row). The area of the bubble is proportional to the magnitude of the corresponding standardised residuals. For positive residuals the bubbles are blue/pink, whereas for negative residuals the bubbles are white.


Figure 9: Commercial and survey selectivities-at-length for scenario 7 ( $\sigma_{R}=1.5$, flat survey selectivity from $\mathbf{3 0} \mathrm{cm}$ ).


Figure 10: Catch trajectories by species, and estimated S. fasciatus proportion in the catch for scenarios 2a, 8a and 8b.

APPENDIX A - Data
Note: Units are throughout cm for length and yr for time.

Table A1: Total catch in kt of redfish (all species combined) in management Units 1 and 2.

| Year | Unit 1 | Unit 2 | Year | Unit 1 | Unit 2 | Year | Unit 1 | Unit 2 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1960 | 12.83 | 23.29 | 1980 | 15.54 | 17.13 | 2000 | 1.12 | 10.29 |
| 1961 | 11.06 | 18.33 | 1981 | 22.05 | 21.75 | 2001 | 1.17 | 8.41 |
| 1962 | 7.15 | 21.30 | 1982 | 26.73 | 17.03 | 2002 | 1.22 | 6.45 |
| 1963 | 20.82 | 22.29 | 1983 | 24.97 | 13.47 | 2003 | 0.84 | 7.47 |
| 1964 | 30.52 | 23.19 | 1984 | 35.83 | 8.14 | 2004 | 0.94 | 5.89 |
| 1965 | 52.83 | 21.83 | 1985 | 28.33 | 11.49 | 2005 | 0.98 | 6.41 |
| 1966 | 67.96 | 28.39 | 1986 | 36.40 | 10.77 | 2006 | 0.69 | 6.48 |
| 1967 | 71.91 | 42.17 | 1987 | 43.45 | 13.96 | 2007 | 0.11 | 3.74 |
| 1968 | 95.26 | 20.17 | 1988 | 51.89 | 10.73 | 2008 | 0.42 | 3.72 |
| 1969 | 92.32 | 46.28 | 1989 | 52.48 | 15.39 | 2009 | 0.60 | 5.13 |
| 1970 | 90.50 | 49.41 | 1990 | 59.90 | 14.79 |  |  |  |
| 1971 | 82.19 | 58.20 | 1991 | 67.53 | 23.21 |  |  |  |
| 1972 | 82.55 | 45.20 | 1992 | 77.75 | 17.16 |  |  |  |
| 1973 | 136.10 | 31.83 | 1993 | 51.09 | 27.43 |  |  |  |
| 1974 | 67.08 | 34.04 | 1994 | 19.39 | 24.32 |  |  |  |
| 1975 | 70.05 | 38.47 | 1995 | 0.05 | 12.24 |  |  |  |
| 1976 | 44.38 | 23.71 | 1996 | 0.07 | 9.41 |  |  |  |
| 1977 | 17.07 | 28.75 | 1997 | 0.04 | 9.94 |  |  |  |
| 1978 | 14.93 | 26.55 | 1998 | 0.40 | 10.64 |  |  |  |
| 1979 | 16.43 | 18.77 | 1999 | 1.11 | 17.90 |  |  |  |

Table A2: Swept area assumed mature (i.e. $>24 \mathrm{~cm}$ for $S$. mentella, and $>22 \mathrm{~cm}$ for $S$. fasciatus) biomass estimates (in kt) and coefficients of variation (CVs) for $S$. mentella and $S$. fasciatus in Units 1 and 2, from McAllister and Duplisea (2012), Table 4.

|  | S. mentella |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
| Year | Unit 1 | CV | Unit 2 | CV |
| 1990 | 443.012 | 0.272 | - | - |
| 1991 | 208.702 | 0.209 | - | - |
| 1992 | 147.726 | 0.206 | - | - |
| 1993 | 93.656 | 0.370 | - | - |
| 1994 | 55.785 | 0.185 | - | - |
| 1995 | 73.626 | 0.112 | - | - |
| 1996 | 59.242 | 0.175 | - | - |
| 1997 | 52.723 | 0.131 | - | - |
| 1998 | 26.391 | 0.186 | - | - |
| 1999 | 47.859 | 0.235 | - | - |
| 2000 | 49.549 | 0.122 | 223.464 | 0.233 |
| 2001 | 43.549 | 0.139 | 151.356 | 0.140 |
| 2002 | 67.468 | 0.797 | - | - |
| 2003 | 95.821 | 0.609 | 100.795 | 0.196 |
| 2004 | 23.963 | 0.219 | - | - |
| 2005 | 46.166 | 0.106 | 90.993 | 0.118 |
| 2006 | 25.042 | 0.125 | - | - |
| 2007 | 28.034 | 0.094 | 76.633 | 0.185 |
| 2008 | 79.371 | 0.462 | - | - |
| 2009 | 11.550 | 0.147 | 103.860 | 0.164 |


| S. fasciatus |  |  |  |
| :---: | :---: | :---: | :---: |
| Unit 1 | CV | Unit 2 | CV |
| 267.287 | - | - | - |
| 188.551 | - | - | - |
| 208.862 | - | - | - |
| 108.936 | - | - | - |
| 70.997 | - | - | - |
| 11.269 | - | - | - |
| 10.183 | - | - | - |
| 26.261 | - | - | - |
| 47.989 | - | - | - |
| 13.266 | - | - | - |
| 19.033 | - | 119.324 | 0.498 |
| 21.572 | - | 177.111 | 0.7 |
| 13.495 | - | - | - |
| 71.947 | - | 69.214 | 0.144 |
| 14.234 | - | - | - |
| 24.429 | - | 168.187 | 0.277 |
| 37.737 | - | - | - |
| 24.09 | - | 158.346 | 0.145 |
| 52.778 | - | - | - |
| 18.683 | - | 127.709 | 0.694 |

Table A3a: Commercial catch-at-length (number) for Atlantic redfish (both species combined) in Unit 1 (Daniel Duplisea, pers. commn)

| Length | 1981 | 1982 | 1983 | 1984 | 1985 | 1986 | 1987 | 1988 | 1989 | 1990 | 1991 | 1992 | 1993 | 1994 | 1999 | 2000 | 2001 | 2002 | 2003 | 2004 | 2005 | 2006 | 2007 | 2008 | 2009 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 10- | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 11 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 12 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 13 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | - | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 14 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 15 | 0 | 0 | 0 | 25 | 9 | 5 | 34 | 24 | 4 | 18 | 5 | 20 | 69 | 8 | 3 | 0 | 1 | 2 | 1 | 0 | 2 | 0 | 0 | 0 | 0 |
| 16 | 0 | 5 | 12 | 78 | 15 | 85 | 23 | 11 | 4 | 33 | 56 | 108 | 1455 | 39 | 5 | 0 | 2 | 6 | 4 | 4 | 1 | 0 | 0 | 0 | 0 |
| 17 | 0 | 0 | 1 | 60 | 47 | 64 | 173 | 24 | 2 | 37 | 82 | 102 | 561 | 28 | 10 | 1 | 4 | 9 | 6 | 5 | 3 | 0 | 0 | 0 | 3 |
| 18 | 3 | 1 | 10 | 42 | 41 | 175 | 356 | 71 | 8 | 41 | 50 | 205 | 504 | 38 | 8 | 1 | 1 | 10 | 7 | 11 | 5 | 4 | 0 | 1 | 1 |
| 19 | 24 | 7 | 1 | 70 | 60 | 169 | 786 | 72 | 5 | 45 | 65 | 307 | 309 | 30 | 10 | 1 | , | 4 | 3 | 11 | 7 | 9 | 0 | 3 | 4 |
| 20 | 75 | 30 | 26 | 272 | 121 | 400 | 1378 | 189 | 30 | 22 | 50 | 313 | 227 | 46 | 14 | 3 | 7 | 5 | , | 14 | 10 | 28 | 0 | 1 | , |
| 21 | 157 | 73 | 78 | 429 | 330 | 790 | 2306 | 518 | 75 | 45 | 113 | 278 | 461 | 34 | 20 | 3 | 11 | 14 | 10 | 7 | 10 | 46 | 0 | 0 | 7 |
| 22 | 170 | 87 | 103 | 372 | 365 | 843 | 3988 | 1700 | 569 | 79 | 154 | 336 | 264 | 58 | 17 | 4 | 11 | 19 | 13 | 4 | 13 | 37 | 0 | 1 | 4 |
| 23 | 228 | 272 | 258 | 395 | 786 | 1232 | 5177 | 4603 | 1815 | 433 | 349 | 438 | 475 | 105 | 21 | 5 | 11 | 26 | 18 | 10 | 18 | 35 | 0 | 2 | 3 |
| 24 | 981 | 434 | 546 | 437 | 1354 | 2300 | 5919 | 10401 | 6025 | 1530 | 957 | 902 | 487 | 215 | 16 | 10 | 21 | 30 | 21 | 16 | 13 | 35 | 0 | 3 |  |
| 25 | 2987 | 1212 | 769 | 810 | 1620 | 3337 | 4300 | 15548 | 13354 | 5457 | 2220 | 1965 | 923 | 461 | 21 | 11 | 16 | 60 | 42 | 29 | 17 | 32 | 0 | 6 | 10 |
| 26 | 6335 | 2301 | 1338 | 1394 | 1600 | 4632 | 3519 | 14592 | 19007 | 15571 | 6771 | 6198 | 2684 | 949 | 24 | 15 | 25 | 50 | 35 | 31 | 22 | 80 | 0 | 5 | 27 |
| 27 | 10618 | 6007 | 2480 | 2286 | 1760 | 5415 | 3505 | 8669 | 19823 | 24636 | 15194 | 14648 | 6809 | 2001 | 37 | 21 | 47 | 60 | 42 | 37 | 42 | 103 | 0 | 8 | 29 |
| 28 | 10985 | 10642 | 5281 | 3829 | 2646 | 5341 | 3770 | 4675 | 13187 | 25363 | 22146 | 22907 | 15034 | 3773 | 51 | 27 | 69 | 66 | 47 | 58 | 45 | 128 | 1 | 16 | 36 |
| 29 | 7815 | 12281 | 8692 | 5891 | 3651 | 5150 | 4037 | 3825 | 7784 | 18290 | 20968 | 25930 | 19200 | 6063 | 86 | 74 | 102 | 50 | 35 | 38 | 40 | 106 | 2 | 18 | 55 |
| 30 | 4720 | 10130 | 9495 | 9479 | 5878 | 6821 | 4835 | 4659 | 6613 | 11038 | 16180 | 21442 | 17271 | 6834 | 192 | 129 | 167 | 69 | 49 | 56 | 63 | 144 | 1 | 27 | 52 |
| 31 | 2534 | 6544 | 8512 | 9733 | 6747 | 7889 | 6239 | 6345 | 6501 | 8279 | 11062 | 14932 | 11961 | 5340 | 216 | 196 | 225 | 132 | 93 | 94 | 69 | 121 | 2 | 34 | 51 |
| 32 | 2214 | 3939 | 6083 | 8760 | 7413 | 8111 | 7989 | 7396 | 7119 | 7951 | 8619 | 10861 | 7465 | 3946 | 282 | 283 | 258 | 185 | 130 | 111 | 88 | 102 | 4 | 36 | 60 |
| 33 | 2007 | 2778 | 3635 | 6919 | 6577 | 7587 | 8202 | 8843 | 7559 | 6839 | 7437 | 9490 | 5367 | 2901 | 252 | 304 | 270 | 227 | 160 | 140 | 122 | 92 | 10 | 37 | 60 |
| 34 | 1553 | 2045 | 2325 | 5168 | 5137 | 5996 | 8427 | 8570 | 6990 | 7107 | 7268 | 9020 | 4971 | 2314 | 244 | 221 | 265 | 256 | 180 | 180 | 139 | 99 | 13 | 48 | 74 |
| 35 | 950 | 1620 | 1803 | 3842 | 3473 | 4298 | 6745 | 7105 | 5347 | 5561 | 5970 | 7577 | 4405 | 2248 | 171 | 220 | 211 | 218 | 153 | 184 | 164 | 68 | 9 | 56 | 82 |
| 36 | 1154 | 1392 | 1437 | 3176 | 2524 | 3129 | 4972 | 4947 | 3997 | 4212 | 4080 | 6475 | 3481 | 1804 | 135 | 163 | 198 | 202 | 142 | 160 | 155 | 71 | 17 | 57 | 68 |
| 37 | 894 | 1286 | 1330 | 2531 | 1998 | 2182 | 3622 | 3794 | 2921 | 3020 | 3277 | 5148 | 3301 | 1070 | 93 | 103 | 114 | 141 | 100 | 136 | 145 | 57 | 19 | 53 | 54 |
| 38 | 743 | 632 | 910 | 2134 | 1783 | 1859 | 2974 | 2754 | 2053 | 2087 | 2367 | 3942 | 2529 | 814 | 70 | 73 | 75 | 100 | 71 | 80 | 114 | 42 | 15 | 47 | 67 |
| 39 | 640 | 445 | 580 | 1723 | 1057 | 1475 | 2051 | 2014 | 1465 | 1627 | 1746 | 3015 | 2124 | 634 | 48 | 49 | 36 | 67 | 47 | 63 | 86 | 25 | 15 | 39 | 46 |
| 40 | 622 | 338 | 403 | 1119 | 822 | 815 | 1489 | 1420 | 1004 | 988 | 1123 | 1977 | 1361 | 486 | 35 | 26 | 30 | 54 | 38 | 40 | 58 | 19 | 8 | 28 | 37 |
| 41 | 524 | 239 | 212 | 535 | 445 | 537 | 879 | 896 | 769 | 518 | 708 | 1334 | 810 | 173 | 20 | 25 | 9 | 39 | 27 | 18 | 33 | 11 | 6 | 23 | 27 |
| 42 | 120 | 133 | 100 | 367 | 353 | 356 | 663 | 561 | 439 | 275 | 390 | 951 | 551 | 118 | 11 | 9 | 3 | 18 | 12 | 10 | 22 | 4 | 4 | 14 | 12 |
| 43 | 25 | 81 | 83 | 114 | 219 | 198 | 323 | 363 | 271 | 200 | 224 | 534 | 295 | 45 | 5 | 13 | 3 | 14 | 10 | 8 | 13 | 3 | 5 | 7 | 7 |
| 44 | 2 | 84 | 46 | 66 | 188 | 127 | 168 | 249 | 119 | 100 | 108 | 320 | 155 | 29 | 2 | 8 | 2 | 9 | 6 | 8 | 10 | 2 | 1 | 7 | 8 |
| 45 | 8 | 72 | 25 | 59 | 58 | 44 | 77 | 91 | 47 | 38 | 73 | 128 | 122 | 12 | 1 | 5 | 1 | 3 | 2 | 1 | 3 | 3 | 2 | 7 | 3 |
| 46 | 0 | 54 | 37 | 28 | 23 | 53 | 47 | 43 | 27 | 15 | 33 | 76 | 49 | 8 | 0 | 3 | 1 | 1 | 1 | 2 | 2 | 1 | 1 | 2 | 3 |
| 47 | 8 | 89 | 51 | 12 | 20 | 26 | 28 | 26 | 9 | 15 | 12 | 29 | 13 | 5 | 0 | 1 | 0 | 2 | 1 | 2 | 2 | , | 1 | 5 | 1 |
| 48 | 1 | 81 | 31 | 7 | 11 | 7 | 23 | 26 | 1 | 2 | 2 | 15 | 3 | 0 | 0 | 2 | 1 | 1 | 1 | 1 | 1 | 0 | 0 | 2 | 1 |
| 49 | 1 | 67 | 43 | 10 | 16 | 4 | 1 | 6 | 5 | 0 | 0 | 0 | 1 | 2 | 0 | 1 | 0 | 1 | 1 | 0 | 0 | 0 | 0 | 2 | 0 |
| 50 | 0 | 95 | 13 | 14 | 14 | 2 | 6 | 1 | 0 | 16 | 0 | 6 | 8 | 0 | 0 | 1 | 3 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 |
| 51 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 52 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 53 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 54 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 55+ | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |

Table A3b: Commercial catch-at-length (numbers) for Atlantic redfish (both species combined) for Unit 2 (Don Power, pers. commn)

| Length | 1995 | 1996 | 1997 | 1998 | 1999 | 2000 | 2001 | 2002 | 2003 | 2004 | 2009 |
| :---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| $10-$ | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 11 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 12 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 13 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 14 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 3 | 0 | 0 |
| 15 | 6 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 2 | 0 | 0 |
| 16 | 13 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 2 | 0 | 1 |
| 17 | 45 | 0 | 0 | 8 | 2 | 1 | 3 | 0 | 6 | 0 | 10 |
| 18 | 148 | 0 | 0 | 0 | 0 | 4 | 5 | 2 | 10 | 15 | 24 |
| 19 | 389 | 0 | 0 | 17 | 4 | 6 | 13 | 4 | 12 | 6 | 39 |
| 20 | 458 | 1 | 0 | 0 | 3 | 5 | 47 | 15 | 31 | 0 | 39 |
| 21 | 521 | 2 | 111 | 18 | 2 | 3 | 41 | 43 | 69 | 31 | 51 |
| 22 | 1104 | 1 | 259 | 17 | 14 | 9 | 101 | 65 | 100 | 52 | 22 |
| 23 | 1489 | 3 | 444 | 38 | 25 | 17 | 136 | 98 | 142 | 119 | 55 |
| 24 | 1123 | 5 | 628 | 49 | 50 | 14 | 356 | 129 | 232 | 156 | 141 |
| 25 | 1279 | 3 | 924 | 157 | 97 | 15 | 521 | 178 | 342 | 187 | 243 |
| 26 | 1708 | 3 | 483 | 273 | 132 | 17 | 745 | 236 | 445 | 264 | 519 |
| 27 | 1966 | 55 | 667 | 346 | 156 | 31 | 640 | 344 | 530 | 330 | 660 |
| 28 | 2592 | 323 | 739 | 487 | 226 | 78 | 643 | 343 | 531 | 267 | 923 |
| 29 | 3191 | 1266 | 1059 | 1059 | 593 | 212 | 565 | 298 | 543 | 302 | 944 |
| 30 | 3364 | 2321 | 1366 | 1793 | 1127 | 425 | 576 | 454 | 636 | 376 | 1064 |
| 31 | 3434 | 2756 | 1435 | 2471 | 1918 | 731 | 751 | 529 | 787 | 473 | 1001 |
| 32 | 2746 | 2817 | 1995 | 2886 | 2455 | 1138 | 914 | 632 | 1098 | 882 | 1082 |
| 33 | 1733 | 2106 | 1779 | 2562 | 2234 | 1244 | 1063 | 730 | 1299 | 1168 | 1007 |
| 34 | 1282 | 1421 | 1780 | 1958 | 2113 | 1100 | 998 | 657 | 1414 | 1405 | 1080 |
| 35 | 842 | 1199 | 1527 | 1599 | 1414 | 851 | 879 | 501 | 1257 | 1330 | 813 |
| 36 | 649 | 855 | 1063 | 1036 | 924 | 592 | 704 | 475 | 1053 | 1184 | 726 |
| 37 | 410 | 676 | 852 | 831 | 619 | 359 | 467 | 328 | 842 | 888 | 576 |
| 38 | 281 | 515 | 543 | 672 | 467 | 306 | 296 | 196 | 499 | 561 | 401 |
| 39 | 212 | 428 | 652 | 462 | 384 | 219 | 214 | 130 | 300 | 405 | 395 |
| 40 | 198 | 320 | 268 | 342 | 252 | 129 | 155 | 94 | 170 | 116 | 170 |
| 41 | 106 | 214 | 324 | 198 | 179 | 75 | 90 | 55 | 106 | 93 | 108 |
| 42 | 66 | 141 | 131 | 107 | 93 | 53 | 94 | 51 | 83 | 33 | 30 |
| 43 | 41 | 90 | 106 | 73 | 63 | 24 | 41 | 40 | 79 | 22 | 16 |
| 44 | 34 | 41 | 82 | 32 | 38 | 18 | 30 | 31 | 58 | 9 | 6 |
| 45 | 18 | 25 | 38 | 16 | 20 | 3 | 23 | 26 | 55 | 5 | 2 |
| 46 | 13 | 6 | 35 | 7 | 6 | 4 | 11 | 18 | 39 | 6 | 4 |
| 47 | 8 | 8 | 0 | 3 | 1 | 1 | 8 | 19 | 34 | 2 | 0 |
| 48 | 0 | 2 | 1 | 2 | 0 | 0 | 0 | 8 | 23 | 0 | 1 |
| 49 | 0 | 0 | 1 | 0 | 0 | 0 | 5 | 4 | 14 | 0 | 0 |
| 50 | 7 | 0 | 0 | 0 | 0 | 0 | 5 | 2 | 14 | 0 | 1 |
| 51 | 0 | 0 | 0 | 1 | 0 | 0 | 2 | 1 | 6 | 0 | 0 |
| 52 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 1 | 10 | 0 | 0 |
| 53 | 1 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 5 | 0 | 0 |
| 54 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 4 | 0 | 0 |
| $55+$ | 0 | 0 | 0 | 0 | 0 | 0 | 8 | 0 | 4 | 0 | 0 |
|  |  |  |  |  |  |  |  |  |  |  |  |
|  | 0 | 0 | 0 |  |  |  |  |  |  |  |  |

Table A4a: Survey catch-at-length (numbers) for S. mentella for Unit 1 and Unit 2 (Daniel Duplisea, pers. commn)


Table A4b: Survey catch-at-length (numbers) for S. fasciatus for Unit 1 and Unit 2 (Daniel Duplisea, pers. commn)

|  | Unit 1 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | Unit 2 |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| length | 1990 | 1991 | 1992 | 1993 | 1994 | 1995 | 1996 | 1997 | 1998 | 1999 | 000 | 2001 | 2002 | 003 | 2004 | 2005 | 2006 | 007 | 2008 | 2009 | 2010 | 2000 | 201 | 2003 | 2005 | 2007 | 2009 |
| 1 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 |
| 2 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 |
| 3 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 |
| 4 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.013 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | .00 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 |
| 5 | 0.011 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.008 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.002 | 0.017 | 0.132 | 0.002 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 |
| 6 | 0.020 | 0.033 | 0.061 | 0.000 | 0.000 | 0.017 | 0.019 | 0.000 | 0.219 | 0.007 | 0.002 | 0.000 | 0.000 | 0.000 | 0.028 | 0.009 | 0.000 | ${ }^{0.310}$ | 0.036 | 0.014 | 0.000 | ${ }^{0} .000$ | 0.000 | 0.000 | 0.187 | 0.000 | 0.000 |
| 7 | 4.208 | 0.576 | 0.600 | 0.075 | 0.000 | 0.275 | 0.360 | 0.468 | 3.187 | 0.053 | 0.313 | 0.282 | 0.010 | 0.027 | 0.174 | 0.735 | 0.010 | 0.125 | 0.479 | 0.745 | 0.192 | 0.060 | 0.000 | 0.010 | 3.499 | 0.152 | 0.080 |
|  | 19.052 | 4.755 | 0.668 | 0.205 | 0.018 | 0.952 | 1.093 | 0.855 | 12.224 | 0.279 | 3.789 | 1.488 | 0.038 | 0.412 | 0.395 | 126.070 | 0.325 | 4.052 | 1.702 | 5.668 | 1.338 | 0.071 | 0.290 | 0.391 | 61.827 | 0.583 | 0.429 |
|  | 89.666 | 9.383 | 1.123 | 0.467 | 0.080 | 2.150 | 1.650 | 0.729 | 8.141 | 1.212 | 7.785 | 6.190 | 0.328 | 0.375 | 0.674 | 552.076 | 1.819 | 42.487 | 2.351 | 10.414 | 2.482 | 0.072 | 1.470 | 0.334 | 140.467 | 0.911 | 0.767 |
| 10 | 63.496 | 66.314 | 1.364 | 0.354 | 0.294 | 0.781 | 1.085 | 1.174 | 3.517 | 2.777 | 4.015 | 11.232 | 1.152 | 0.533 | 0.900 | 192.448 | 9.015 | 11.445 | 3.149 | 6.285 | 2.596 | 0.331 | 6.210 | 0.169 | 74.065 | 1.102 | 0.746 |
| 11 | 1.562 | 206.499 | 6.053 | 0.320 | 0.380 | 0.276 | 0.808 | 2.295 | 1.831 | 5.628 | 1.740 | 4.151 | 1.990 | 1.009 | 0.549 | 11.096 | 24.348 | 0.766 | 6.376 | 0.846 | 3.757 | 0.820 | 6.069 | 0.232 | 12.727 | 0.435 | 0.788 |
| 12 | 1.377 | 355.845 | 21.390 | 0.518 | 0.912 | 0.435 | 0.796 | 2.851 | 1.701 | 7.587 | 3.656 | 1.563 | 3.618 | 2.620 | 1.198 | 0.933 | 119.218 | 3.765 | 28.003 | 1.256 | 9.753 | 0.839 | 1.842 | 0.997 | 6.901 | 9.796 | 1.784 |
| 13 | 2.370 | 179.842 | 41.364 | 0.955 | 0.998 | 0.529 | 0.855 | 2.277 | 2.011 | 6.309 | 6.902 | 2.235 | 6.596 | 4.293 | 1.737 | 1.055 | 176.801 | 29.379 | 16.814 | 2.390 | 10.983 | 0.985 | 2.900 | 1.598 | 5.992 | 59.957 | 4.698 |
| 14 | 3.969 | 20.317 | 42.606 | 2.461 | 1.192 | 0.589 | 0.526 | 1.549 | 2.186 | 4.678 | 10.968 | 2.063 | 5.311 | 8.878 | 2.322 | 2.119 | 46.190 | 81.378 | 6.749 | 3.625 | 5.390 | 1.714 | 3.256 | 2.563 | 7.096 | 182.403 | 9.255 |
| 15 | 7.191 | 7.285 | 19.065 | 5.797 | 2.055 | 0.766 | 0.517 | 0.958 | 2.961 | 3.155 | 10.896 | 1.818 | 1.586 | 10.177 | 3.291 | 1.777 | 7.487 | 130.437 | 5.685 | 5.789 | 2.875 | 3.055 | 4.895 | 3.844 | 6.889 | 240.143 | 14.248 |
| 16 | 9.977 | 7.241 | 1.347 | 8.428 | 2.467 | 0.698 | 0.427 | 0.616 | 2.381 | 2.369 | 4.770 | 2.364 | 1.000 | 8.084 | 3.315 | 1.793 | 1.612 | 70.727 | 10.320 | 6.111 | 2.951 | 3.437 | 6.823 | 4.002 | 10.074 | 120.990 | 16.145 |
| 17 | 14.364 | 7.989 | 1.262 | 6.582 | 2.539 | 0.927 | 0.462 | 0.450 | 1.327 | 1.755 | 3.346 | 2.148 | 1.181 | 4.784 | 3.530 | 1.420 | 0.552 | 19.580 | 10.806 | 2.572 | 2.959 | 5.827 | 10.228 | 4.435 | 22.083 | 37.332 | 46.546 |
| 1 | 1122 | 6565 | 1778 | ${ }^{453}$ | ) $\quad$ n) | ก 70 | $\bigcirc 450$ | ก 424 | 1 na9 | 1157 | 2710 | 1425 | 1051 | ${ }^{2} 118$ | ${ }^{3} 640$ | 1875 | 1150 | 3756 | ${ }^{\circ} 988$ | ${ }^{3} 188$ | 3149 | 7767 | $11^{158}$ | 5130 | 37597 | 15951 | 84143 |
| 19 | 3.876 | 4.305 | 1.217 | 0.856 | 0.893 | 0.552 | 0.694 | 0.469 | 0.846 | 0.778 | 1.246 | 1.011 | 0.840 | 1.465 | 2.174 | 2.044 | 1.229 | 1.907 | 3.692 | 3.360 | 2.487 | 9.533 | 11.138 | 6.331 | 50.165 | 10.533 | 83.373 |
| 20 | 1.582 | 2.148 | 1.120 | 0.600 | 0.440 | 0.500 | 0.560 | 0.450 | 1.636 | 0.401 | 1.009 | 0.694 | 0.879 | 1.103 | 1.263 | 2.018 | 1.348 | 1.752 | 1.369 | 2.338 | 2.149 | 9.798 | 10.626 | 8.022 | 50.734 | 13.346 | 59.069 |
| 21 | 1.222 | 1.963 | 1.313 | 0.813 | 0.185 | 0.367 | 0.630 | 0.366 | 1.406 | 0.346 | 0.390 | 0.559 | 0.697 | 0.964 | 0.596 | 1.365 | 1.422 | 2.194 | 0.635 | 1.716 | 1.516 | 8.069 | 10.094 | 10.871 | 37.204 | 10.305 | 29.014 |
| 22 | 1.524 | 1.307 | 1.810 | 2.039 | 0.219 | 0.356 | 0.376 | 0.352 | 4.929 | 0.328 | 0.582 | 0.582 | 0.685 | 1.039 | 0.563 | 1.006 | 1.468 | 1.044 | 0.521 | 1.205 | 1.321 | 6.802 | 7.924 | 13.986 | 27.164 | 11.562 | 11.604 |
| 23 | 1.753 | 1.631 | 3.170 | 4.818 | 0.389 | 0.264 | 0.239 | 0.251 | 3.871 | 0.447 | 0.310 | 0.336 | 0.407 | 0.965 | 0.612 | 0.594 | 2.151 | 0.776 | 0.544 | 0.664 | 0.862 | 6.001 | 10.150 | 10.622 | 19.816 | 12.633 | 5.769 |
| 24 | 3.181 | 2.298 | 4.075 | 8.224 | 0.603 | 0.250 | 0.185 | 0.347 | 5.376 | 0.381 | 0.440 | 0.333 | 0.350 | 1.454 | 0.781 | 0.453 | 1.629 | 0.802 | 0.823 | 0.448 | 0.418 | 7.882 | 25.295 | 9.675 | 18.605 | 12.181 | 15.870 |
| 25 | 6.559 | 3.464 | 4.070 | 7.765 | 0.764 | 0.346 | 0.130 | 0.264 | 3.136 | 0.336 | 0.321 | 0.307 | 0.556 | 1.779 | 0.813 | 0.453 | 2.209 | 0.480 | 0.935 | 0.536 | 0.374 | 9.976 | 37.601 | 8.813 | 16.561 | 16.012 | 20.152 |
| 26 | 13.683 | 5.013 | 5.560 | 7.992 | 1.508 | 0.299 | 0.183 | 0.284 | 2.974 | 0.374 | 0.221 | 0.376 | 0.286 | 2.750 | 0.330 | 0.658 | 2.851 | 0.287 | 1.383 | 0.746 | 0.660 | 11.383 | 65.737 | 10.033 | 15.435 | 19.007 | 17.919 |
| 27 | 22.599 | 9.103 | 9.703 | 9.571 | 2.167 | 0.237 | 0.152 | 0.343 | 2.477 | 0.487 | 0.278 | 0.294 | 0.500 | 3.749 | 0.984 | 0.764 | 2.432 | 0.507 | 1.244 | 0.675 | 0.739 | 10.200 | 47.704 | 8.738 | 12.501 | 17.462 | 16.557 |
| 28 | 28.886 | 13.078 | 14.215 | 7.937 | 1.545 | 0.233 | 0.159 | 0.703 | 1.298 | 0.424 | 0.213 | 0.202 | 0.383 | 5.810 | 0.628 | 0.630 | 1.956 | 0.431 | 1.260 | 0.626 | 0.807 | 8.029 | 32.294 | 7.496 | 8.120 | 13.448 | 14.291 |
| 29 | 22.941 | 15.507 | 14.714 | 5.745 | 2.436 | 0.345 | 0.406 | 0.930 | 2.401 | 0.437 | 0.346 | 0.295 | 0.398 | 7.156 | 0.796 | 0.582 | 1.638 | 0.451 | 1.489 | 0.773 | 0.915 | 7.236 | 23.948 | 7.172 | 4.922 | 7.557 | 10.712 |
| 30 | 13.174 | 12.140 | 12.670 | 6.036 | 3.072 | 0.300 | 0.492 | 1.216 | 2.331 | 0.421 | 0.473 | 0.314 | 0.441 | 5.158 | 0.565 | 0.549 | 1.414 | 0.341 | 2.175 | 0.610 | 0.387 | 7.494 | 26.153 | 6.663 | 5.574 | 8.138 | 9.081 |
| 31 | 7.520 | 8.361 | 9.134 | 4.958 | 2.319 | 0.348 | 0.404 | 1.464 | 1.920 | 0.276 | 0.446 | 0.665 | 0.370 | 1.908 | 0.517 | 0.558 | 0.856 | 0.265 | 1.915 | 0.624 | 0.496 | 7.481 | 10.925 | 5.396 | 4.168 | 6.666 | 8.268 |
| 32 | 4.622 | 5.607 | 8.374 | 2.606 | 2.708 | 0.258 | 0.380 | 1.212 | 0.572 | 0.307 | 0.510 | 0.826 | 0.463 | 2.306 | 0.219 | 0.573 | 0.731 | 0.255 | 2.491 | 0.485 | 0.397 | 8.830 | 9.416 | 4.748 | 5.852 | 7.385 | 7.008 |
| 33 | 3.425 | 3.643 | 4.935 | 1.636 | 2.397 | 0.195 | 0.310 | 1.084 | 0.666 | 0.358 | 0.661 | 0.885 | 0.258 | 0.802 | 0.156 | 0.511 | 0.538 | 0.335 | 2.395 | 0.386 | 0.214 | 7.006 | 3.172 | 2.908 | 5.949 | 7.342 | 5.370 |
| 34 | 4.006 | 2.716 | 3.766 | 0.963 | 1.866 | 0.230 | 0.196 | 0.887 | 0.484 | 0.373 | 0.505 | 0.695 | 0.311 | 0.685 | 0.051 | 0.450 | 0.439 | 0.351 | 1.154 | 0.319 | 0.379 | 7.938 | 2.791 | 3.133 | 6.745 | 6.537 | 5.569 |
| 35 | 3.331 | 2.503 | 3.208 | 0.620 | 1.478 | 0.280 | 0.220 | 0.821 | 0.808 | 0.313 | 0.465 | 0.700 | 0.342 | 0.459 | 0.105 | 0.509 | 0.884 | 0.381 | 1.572 | 0.204 | 0.578 | 8.769 | 1.635 | 2.809 | 6.516 | 4.566 | 5.133 |
| 36 | 3.614 | 2.241 | 1.655 | 0.342 | 1.425 | 0.206 | 0.175 | 0.418 | 0.291 | 0.283 | 0.524 | 0.476 | 0.311 | 0.522 | 0.032 | 0.884 | 0.330 | 0.424 | 1.044 | 0.198 | 0.507 | 5.125 | 1.509 | 2.184 | 5.120 | 3.688 | 4.370 |
| 37 | 2.555 | 1.655 | 2.130 | 0.312 | 1.180 | 0.172 | 0.137 | 0.198 | 0.228 | 0.290 | 0.363 | 0.591 | 0.202 | 0.469 | 0.091 | 0.889 | 0.341 | 0.318 | 0.748 | 0.238 | 0.436 | 5.339 | 1.077 | 1.522 | 4.309 | 3.297 | 4.452 |
| 38 | 2.357 | 1.749 | 0.907 | 0.162 | 1.056 | 0.200 | 0.134 | 0.110 | 0.196 | 0.092 | 0.351 | 0.310 | 0.132 | 0.677 | 0.047 | 0.281 | 0.279 | 0.281 | 0.497 | 0.145 | 0.197 | 3.786 | 0.376 | 1.311 | 3.195 | 2.842 | 3.351 |
| 39 | 1.990 | 1.188 | 1.056 | 0.072 | 0.771 | 0.097 | 0.097 | 0.124 | 0.160 | 0.101 | 0.182 | 0.165 | 0.146 | 0.527 | 0.027 | 0.143 | 0.159 | 0.662 | 0.429 | ${ }^{0.157}$ | 0.179 | 2.300 | 0.244 | 0.943 | 1.678 | 1.846 | 2.361 |
| 40 | 1.165 | 0.970 | 0.675 | 0.054 | 0.414 | 0.100 | 0.074 | 0.121 | 0.091 | 0.079 | 0.152 | 0.158 | 0.055 | 0.409 | 0.023 | 0.178 | 0.169 | 0.148 | 0.417 | 0.135 | 0.141 | 1.961 | 0.242 | 0.640 | 1.219 | 1.615 | 1.535 |
| 41 | 1.051 | 0.717 | 0.278 | 0.041 | 0.183 | 0.098 | 0.042 | 0.029 | 0.054 | 0.049 | 0.116 | 0.114 | 0.057 | 0.151 | 0.036 | 0.137 | 0.112 | 0.103 | 0.250 | 0.091 | 0.226 | 0.867 | 0.122 | 0.383 | 0.603 | 0.715 | 0.805 |
| 42 | 0.500 | 0.381 | 0.180 | 0.041 | 0.084 | 0.065 | 0.027 | 0.035 | 0.041 | 0.031 | 0.067 | 0.045 | 0.025 | 0.113 | 0.067 | 0.064 | 0.086 | 0.074 | 0.110 | 0.061 | 0.091 | 1.301 | 0.089 | 0.278 | 0.581 | 0.619 | 0.427 |
| 43 | 0.322 | 0.224 | 0.105 | 0.008 | 0.096 | 0.018 | 0.020 | 0.045 | 0.029 | 0.021 | 0.056 | 0.048 | 0.006 | 0.026 | 0.028 | 0.032 | 0.052 | 0.046 | 0.108 | 0.054 | 0.056 | 1.047 | 0.037 | 0.162 | 0.223 | 0.302 | 0.268 |
| 44 | 0.242 | 0.190 | 0.055 | 0.016 | 0.038 | 0.029 | 0.012 | 0.017 | 0.027 | 0.013 | 0.037 | 0.005 | 0.007 | 0.033 | 0.006 | 0.082 | 0.016 | 0.021 | 0.037 | 0.019 | 0.027 | 0.790 | 0.038 | 0.135 | 0.154 | 0.187 | 0.168 |
| 45 | 0.095 | 0.115 | 0.021 | 0.001 | 0.017 | 0.007 | 0.002 | 0.014 | 0.013 | 0.022 | 0.018 | 0.029 | 0.008 | 0.024 | 0.017 | 0.037 | 0.012 | 0.121 | 0.020 | 0.012 | 0.012 | 0.357 | 0.005 | 0.107 | 0.092 | 0.083 | 0.122 |
| 46 | ${ }^{0.037}$ | 0.054 | 0.029 | 0.000 | 0.013 | 0.008 | 0.000 | 0.002 | 0.000 | 0.005 | 0.005 | 0.004 | 0.000 | 0.004 | 0.004 | 0.017 | 0.050 | 0.019 | 0.024 | 0.008 | 0.012 | 0.072 | 0.009 | 0.059 | 0.015 | 0.075 | 0.038 |
| 47 | 0.011 | 0.026 | 0.012 | 0.000 | 0.004 | 0.002 | 0.008 | 0.000 | 0.007 | 0.001 | 0.006 | 0.005 | 0.000 | 0.000 | 0.002 | 0.049 | 0.000 | 0.000 | 0.003 | 0.000 | 0.014 | 0.073 | 0.006 | 0.042 | 0.032 | 0.022 | 0.041 |
| 48 | ${ }^{0.006}$ | ${ }^{0.006}$ | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.006 | ${ }^{0.001}$ | 0.002 | 0.002 | 0.000 | 0.027 | 0.002 | 0.139 | 0.016 | ${ }^{0.003}$ | 0.000 | 0.002 | 0.011 | ${ }^{0.141}$ | 0.010 | 0.043 | 0.021 | ${ }^{0.008}$ | 0.071 |
| 49 | 0.007 | 0.020 | 0.001 | 0.000 | 0.007 | 0.000 | 0.000 | 0.013 | 0.008 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.002 | 0.045 | 0.000 | 0.000 | 0.001 | 0.005 | 0.009 | 0.033 | 0.009 | 0.035 | 0.007 | 0.003 | ${ }_{0} 0.006$ |
| 50 | 0.001 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.004 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.002 | 0.049 | 0.000 | 0.007 | 0.066 | 0.000 | 0.000 | 0.000 | 0.007 | 0.032 | 0.000 | 0.000 | 0.011 |
| 51 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.007 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.002 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.012 | 0.015 | 0.021 | 0.000 | 0.010 | 0.012 |
| 52 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.001 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.039 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.024 | 0.002 | 0.016 | 0.000 | 0.000 | 0.000 |
| 53 | 0.015 | 0.000 | 0.000 | 0.000 | 0.007 | 0.000 | 0.000 | 0.000 | 0.004 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.012 | 0.013 | 0.012 | 0.000 | 0.000 | 0.000 |
| 54 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.019 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.013 | 0.000 | 0.000 | 0.000 |
| 55 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.010 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.004 | 0.000 | 0.000 | 0.000 |
| 56 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.002 | 0.010 | 0.000 | 0.000 | 0.000 | 0.000 | 0.001 | 0.000 | 0.001 | 0.000 | 0.000 | 0.000 |  |
| 57 | 0.000 | 0.015 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 |
| 58 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.00 |
| 59 | 0.000 | .00 | 0.000 | 0.000 | 00 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.00 | 0.000 | 0.000 | 0.000 | 0.00 | 0.001 | 0.000 | 0.000 | 0.000 | 0.00 |

Table A5: Life history parameters assumed for S. mentella and S. fasciatus.

| S. mentella |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
| M | 0.1 |  |  | McAllister and Duplisea (2012) |
| $h$ | 0.67 |  |  | McAllister and Duplisea (2012) |
| Length-at-maturity | 24 |  |  | Knife-edged, Don Power, pers. commn |
| Fraction of $M$ that occurs before spawning ( $M^{5}$ ) | 0.25 |  |  |  |
|  | $L_{\text {inf }}$ | $\kappa$ | $t_{0}$ |  |
| Length-at-age | 35.81 | 0.1458 | 0 | $L_{a}=L_{\text {inf }}\left(1-e^{-\kappa\left(a-t_{0}\right)}\right)$, Campana, pers. commn |
|  | $\alpha$ | $\beta$ |  |  |
| Weight-at-age | 0.00944 | 3.107 |  | $W_{a}=\alpha\left(L_{a}\right)^{\beta} \quad, \mathrm{McAllister}$ and Duplisea (2012) |
| S. fasciatus |  |  |  |  |
| M | 0.125 |  |  | McAllister and Duplisea (2012) |
| $h$ | 0.67 |  |  | McAllister and Duplisea (2012) |
| Length-at-maturity | 22 |  |  | Knife-edged, Don Power, pers. commn |
| Fraction of $M$ that occurs before spawning ( $M^{5}$ ) |  |  |  |  |
| Length-at-age | $L_{\text {inf }}$ | $\kappa$ | $t_{0}$ |  |
|  | 31.88 | 0.2213 | 0 | $L_{a}=L_{\text {inf }}\left(1-e^{-\kappa\left(\alpha-t_{0}\right)}\right)$, Campana, pers. commn |
|  | $\alpha$ | $\beta$ |  |  |
| Weight-at-age | 0.01106 | 3.08 |  | $W_{a}=\alpha\left(L_{a}\right)^{\beta} \quad, \mathrm{Mc}$ Allister and Duplisea (2012) |

## Appendix B - The Statistical Catch-At-Length Model

The model used for these assessments is a Statistical Catch-At-Length (SCAL) model. The approach used involves the construction of an age-structured model of the population dynamics and fitting it to the available abundance indices by maximising the likelihood function. The general specifications of the model and its equations are described below, followed by details of the contributions to the (penalised) log-likelihood function from the different sources of data available and assumptions concerning the stock-recruitment relationship. Quasi-Newton minimization is used to minimize the total negative loglikelihood function (the package AD Model Builder ${ }^{\top M}$, Otter Research, Ltd is used for this purpose).

## B.1. Population dynamics

## B.1.1 Numbers-at-age

The resource dynamics of the two populations (S. mentella and S. fasciatus) are modelled by the following set of population dynamics equations:

$$
\begin{align*}
& N_{s, y+1,0}=R_{s, y+1}  \tag{B1}\\
& N_{s, y+1, a+1}=N_{s, y, a} e^{-Z_{s, y, a}} \quad \text { for } 0 \leq a \leq m_{s}-2  \tag{B2}\\
& N_{s, y+1, m_{s}}=N_{y, m_{s}-1} e^{-Z_{s, y, m_{s}-1}}+N_{y, m_{s}} e^{-Z_{s, y, m_{s}}}+ \tag{B3}
\end{align*}
$$

where
$N_{s, y, a}$ is the number of species $s$ and age $a$ at the start of year $y$ (which refers to a calendar year),
$R_{s, y} \quad$ is the recruitment (number of 0-year-old fish) of species $s$ at the start of year $y$,
$m_{s} \quad$ is the maximum age considered (taken to be a plus-group) for species $s, m_{s}=20$,
$Z_{s, y, a}=\sum_{u} F_{s, u, y} S_{s, u, y, a}+M_{s, a}$ is the total mortality in year $y$ on fish of species $s$ and age $a$, and
$M_{s, a}$ denotes the natural mortality rate for fish of species $s$ of age $a$,
$F_{s, u, a}$ is the fishing mortality of a fully selected age class of species $s$, for Unit $u$ in year $y$,
$S_{s, y, a}$ is the commercial selectivity (i.e. combination of availability and vulnerability to fishing gear) of species $s$ at age $a$ and in year $y$; when $S_{s, y, a}=1$, the age-class $a$ is said to be fully selected,

Selectivity is estimated as a function of length and then converted to selectivity-at-age:

$$
\begin{equation*}
S_{s, u, y, a}=\sum_{l} S_{s, u, y, l} A_{s, a, l} \tag{B4}
\end{equation*}
$$

where $A_{s, a, l}$ is the proportion of fish of species $s$ and age $a$ that fall in the length group / (i.e., $\sum_{l} A_{s, a, l}=1$ for all ages).

The matrix $A_{s, a, l}$ is calculated under the assumption that length-at-age is normally distributed about a mean given by the von Bertalanffy equation, i.e. (omitting the species subscript s):

$$
\begin{equation*}
L_{a} \sim N\left[L_{\infty}\left(1-e^{-\kappa\left(a-t_{o}\right)}\right) ; \theta_{a}^{2}\right] \tag{B5}
\end{equation*}
$$

where
$\theta_{a}$ is the standard deviation of length-at-age $a$, which is taken as proportional to the expected length-at-age $a$, i.e.:
$\theta_{a}=\beta^{*} L_{\infty}\left(1-e^{-\kappa\left(a-t_{o}\right)}\right)$
with $\beta^{*}$ an estimable parameter (taken to be the same for both species).

## B.1.2. Recruitment

The number of recruits of each species at the start of year $y$ is assumed to be related to the spawning stock size (i.e. the biomass of mature fish) by a Beverton-Holt stock-recruitment relationship (Beverton and Holt, 1957), parameterised in terms of the "steepness" of the stock-recruitment relationship, $h_{s}$, and the pre-exploitation equilibrium spawning biomass, $K_{s}^{s p}$, and recruitment, $R_{s, 0}$ and allowing for annual fluctuation about the deterministic relationship:

$$
\begin{equation*}
R_{s, y}=\frac{4 h_{s} R_{s, 0} B_{s, y}^{s p}}{K_{s}^{s p}\left(1-h_{s}\right)+\left(5 h_{s}-1\right) B_{s, y}^{s p}} e^{\left(\varsigma_{s, y}-\sigma_{R}^{2} / 2\right)} \tag{B7}
\end{equation*}
$$

where
$\zeta_{s, y} \quad$ reflects fluctuation about the expected recruitment for species $s$ for year $y$, which is assumed to be normally distributed with standard deviation $\sigma_{R}$ (which is input in the applications considered here); these residuals are treated as estimable parameters in the model fitting process.
$B_{s, y}^{s p} \quad$ is the spawning biomass of species $s$ at the start of year $y$, computed as:
$B_{s, y}^{s p}=\sum_{a=1}^{m_{s}} f_{s, a} w_{s, a}^{s t r t} N_{s, y, a} e^{-M_{s, a} M_{s}^{s p}}$
where
$w_{s, a}^{s t r t}$ is the mass of fish of species $s$ and age $a$ during spawning,
$f_{s, a}$ is the proportion of fish of species $s$ and age $a$ that are mature
$M_{s}^{s p}$ is the fraction of mortality that occurs before spawning for species $s\left(M_{s}^{s p}=0.25\right)$.
In the fitting procedure, $K_{s}^{s p}$ are estimated while $h_{s}$ have thus far been fixed at 0.67 for consistency with McAllister and Duplisea (2011).

## B.1.3. Total catch and catches-at-age

The catch by mass for species $s$ in Unit $u$, in year $y$ is given by:

$$
\begin{equation*}
C_{s, u, y}=\sum_{a=0}^{m} \tilde{w}_{s, a}^{\text {mid }} C_{s, u, y, a}=\sum_{a=0}^{m} \tilde{w}_{s, y, a}^{\text {mid }} F_{s, u, y} S_{s, u, y, a} N_{s, y, a}\left(1-e^{-Z_{s, y, a}}\right) / Z_{s, y, a} \tag{B9}
\end{equation*}
$$

where
$C_{s, u, y, a}$ is the catch-at-age, i.e. the number of fish of species $s$ and age $a$, caught in year $y$ in Unit $u$,
$\tilde{w}_{s, y, a}^{\text {mid }}$ is the selectivity-weighted mid-year weight-at-age $a$ for species $s$ landed in year $y$, and
$\tilde{w}_{s, y, a}^{\operatorname{mid}}=\sum_{l} S_{s, y, l} w_{s, l} A_{s, a, l} / \sum_{l} S_{s, y, l} A_{s, a, l}$
with
$w_{s, l} \quad$ being the weight of fish of species $s$ and length $/$.

The model estimate of the survey biomass of species $s$ in Unit $u$ is calculated as:

$$
\begin{equation*}
B_{s, u, y}^{s u r v}=\sum_{a=1}^{m} \tilde{w}_{s, y, a}^{\text {mid }} S_{s, u, a}^{s u r v} N_{s, y, a} e^{-Z_{s, y, a} \frac{m_{u}^{s u v}}{12}} \tag{B11}
\end{equation*}
$$

where
$S_{s, u, a}^{s u r v}$ is the survey selectivity for species $s$ and age $a$ in Unit $u$,
$m_{u}^{\text {surv }}$ is the month in which survey takes place in Unit $u\left(m_{u}^{\text {surv }}=8\right)$, and

## B.1.4. Initial conditions

For the first year $\left(y_{0}\right)$ considered in the model therefore, the stock is assumed to be at a fraction $\left(\theta_{s}\right)$ of its pre-exploitation biomass, i.e.:

$$
\begin{equation*}
B_{s, y_{0}}^{s p}=\theta_{s} \cdot K_{s}^{s p} \tag{B12}
\end{equation*}
$$

with the starting age structure:

$$
\begin{equation*}
N_{s, y_{0}, a}=R_{\text {start }, s} N_{\text {start }, s, a} \quad \text { for } 0 \leq a \leq m_{s} \tag{B13}
\end{equation*}
$$

where

$$
\begin{array}{ll}
N_{s t a r t, s, 0}=1 \\
N_{s t a r t, s, a}=N_{s t a r t, s, a-1} e^{-M_{s, a-1}}\left(1-\phi_{s} S_{s, a-1}\right) & \text { for } 1 \leq a \leq m_{s}-1 \\
N_{s t a r t, s, m}=N_{s t a r t, s, m-1} e^{-M_{s, m-1}}\left(1-\phi_{s} S_{s, m_{s}-1}\right) /\left(1-e^{-M_{s, m}}\left(1-\phi_{s} S_{s, m}\right)\right) \tag{B16}
\end{array}
$$

where $\phi_{s}$ characterises the average fishing proportion over the years immediately preceding $y_{0}$.

Unless indicated otherwise though, the stock is assumed to be at pristine equilibrium in 1960, i.e. $\theta_{0}=1$ and $\phi_{s}=0$ for the results reported here.

## B.2. The (penalised) likelihood function

The model is fit to survey abundance indices, and commercial and survey catch-at-length data to estimate model parameters (which may include residuals about the stockrecruitment function, the fishing selectivities, the annual catches or natural mortality,
facilitated through the incorporation of penalty functions described below). Contributions by each of these to the negative of the (penalised) log-likelihood ( $-\ell \mathrm{n} L$ ) are as follows.

## B.2.1. Survey abundance data

The likelihood is calculated assuming that the observed survey index is log-normally distributed about its expected value:
$I_{s, u, y}=\hat{I}_{s, u, y} \exp \left(\varepsilon_{s, y}^{u}\right) \quad$ or $\quad \varepsilon_{s, u, y}=\ell \mathrm{n}\left(I_{s, u, y}\right)-\ell \mathrm{n}\left(\hat{I}_{s, u, y}\right)$
where
$I_{s, u, y}$ is the survey biomass index for year $y$, species $s$ and Unit $u$,
$\hat{I}_{s, u, y}=\hat{q}_{s, u} \hat{B}_{s, u, y}^{s u r v}$ is the corresponding model estimate, where $\widehat{B}_{s, u, y}^{s u r v}$ is the model estimate of survey biomass, given by equation (B10),
$\hat{q}_{s, u} \quad$ is the constant of proportionality (catchability) for species $s$ in Unit $u$, and
$\varepsilon_{s, u, y} \quad$ from $N\left(0,\left(\sigma_{s, u, y}\right)^{2}\right)$.

The contribution of the survey biomass data to the negative of the log-likelihood function (after removal of constants) is then given by:
$-\ln L^{s u r v}=\sum_{s} \sum_{u} \sum_{y}\left\lfloor\ln \left(\sigma_{s, u, y}\right)+\left(\varepsilon_{s, u, y}\right)^{2} / 2\left(\sigma_{s, u, y}\right)^{2}\right]$
where
$\sigma_{s, u, y}$ is the standard deviation of the residuals for the logarithm of survey index for species $s$ in Unit $u$ and year $y$.
The catchability coefficient $q_{s, u}$ for survey index for species $s$ in Unit $u$ is estimated by its maximum likelihood value:
$\ln \hat{q}_{s, u}=1 / n \sum_{y}\left(\ln I_{s, u, y}-\ln \hat{B}_{s, u, y}^{s u r v}\right)$

A penalty on the survey catchability coefficients is used for all scenarios in the spirit of a prior to avoid the results going into implausible regions of parameter space (particularly $S$. fasciatus' survey catchability $q$ going unrealistically high). The following penalty is added to the negative log-likelihood to effect this:

$$
\begin{equation*}
-\ell n L^{q p e n}=\sum_{i=1}^{n_{\text {ancec }}}\left[\frac{2\left(q^{i}-l_{b}\right)}{\left(u_{b}-l_{b}\right)}-1\right]^{p} \tag{B20}
\end{equation*}
$$

## B.2.2. Commercial catches-at-length

Commercial catches-at-length are not disaggregated by species. The model is therefore fit to the catches-at-length as determined for both species combined. The contribution of the catch-at-length data to the negative of the log-likelihood function under the assumption of an "adjusted" (or "Punt-Kennedy (1997)") lognormal error distribution is given by:
$-\ln L^{C A L}=W_{C A L} \sum_{u} \sum_{y} \sum_{l}\left[\ln \left(\sigma_{u}^{c o m} / \sqrt{p_{u, y, l}}\right)+p_{u, y, l}\left(\ln p_{u, y, l}-\ln \hat{p}_{u, y, l}\right)^{2} / 2\left(\sigma_{u}^{c o m}\right)^{2}\right]$
where
$p_{u, y, l}$ is the observed proportion of fish (S. mentella and S. fasciatus combined) caught in year $y$ and Unit $u$ that are of length I,
$\hat{p}_{u, y, l}=\frac{\sum_{s} \hat{C}_{s, u, y, l}}{\sum_{l^{\prime}} \sum_{s} \hat{C}_{s, u, y, l}}$ is the model-predicted proportion of fish (S. mentella and S. fasciatus combined) caught in year $y$ and Unit $u$ that are of length I,
where

$$
\begin{equation*}
\hat{C}_{s, u, y, l}=N_{s, y, a} A_{s, a, l} S_{s, u, y, l} l^{-Z_{s, s, a} / 2} \tag{B22}
\end{equation*}
$$

and $\sigma_{u}^{c o m}$ is the standard deviation associated with the catch-at-length data for Unit $u$, which is estimated in the fitting procedure by:
$\hat{\sigma}_{u}^{c o m}=\sqrt{\sum_{y} \sum_{l} p_{u, y, l}\left(\ln p_{u, y, l}-\ln \hat{p}_{u, y, l}\right)^{2} / \sum_{y} \sum_{l} 1}$

The log-normal error distribution underlying equation (B21) is chosen on the grounds that (assuming no ageing error) variability is likely dominated by a combination of interannual variation in the distribution of fishing effort, and fluctuations (partly as a consequence of such variations) in selectivity-at-age, which suggests that the assumption of a constant coefficient of variation is appropriate. However, for ages poorly represented in the sample, sampling variability considerations must at some stage start to dominate the variance. To take this into account in a simple manner, motivated by binomial distribution properties, the observed proportions are used for weighting so that undue importance is not attached to data based upon a few samples only.

The $W_{C A L}$ weighting factor is set to 0.01 (if not otherwise indicated) to downweight the contribution of the catch-at-length data (which tend to be positively correlated between adjacent length groups) to the overall negative log-likelihood compared to that of the survey biomass data.

Commercial catches-at-length are incorporated in the likelihood function using equation (B20), for which the summation over age $/$ is taken from length $I_{\text {minus }}$ (considered as a minus group) to $I_{\text {plus }}$ (a plus group), see Table B1.

## B.2.3. Survey catches-at-length

The survey catches-at-length are incorporated into the negative of the log-likelihood in an analogous manner to the commercial catches-at-length, assuming an adjusted log-normal error distribution (equation (B20)). In this case however, data that are disaggregated by species are available.
$p_{s, u, y, l}^{s u r v}$ is the observed proportion of fish of species $s$ and length / in year $y$ for survey carried out in Unit $u$,
$\hat{p}_{s, u, y, l}^{\text {surv }}$ is the expected proportion of fish of species $s$ and length $/$ in year $y$ in the survey carried out in Unit $u$, given by:
$\hat{p}_{s, u, y, l}^{s u r v}=\hat{C}_{s, u, l, l}^{s u r v} / \sum_{l} \hat{C}_{s, u, y, l}^{s u r v}$ siters the model-predicted proportion of fish for species $s$ caught in year $y$ and survey carried out in Unit $u$ that are of length I,
where

$$
\begin{equation*}
\hat{C}_{s, s, u, y, l}^{s u r v}=N_{s, y, a} A_{s, a, l}{ }_{s, l}^{s u r v, l} e^{\text {surv }} \tag{B24}
\end{equation*}
$$

Survey catches-at-length are incorporated in the likelihood function using equation (B21), for which the summation over age $I$ is taken from length $I_{\text {minus }}$ (considered as a minus group) to $I_{\text {plus }}$ (a plus group), see Table B1.

## B.2.4. Stock-recruitment function residuals

The stock-recruitment residuals are assumed to be log-normally distributed. Thus, the contribution of the recruitment residuals to the negative of (he (now penalised) loglikelihood function is given by:
$-\ell n L^{S R P e n}=\sum_{s} \sum_{y=y 1}^{y 2}\left[S_{s, y}^{2} / 2 \sigma_{R}^{2}\right]$
where
$\varsigma_{s, y} \quad$ from $N\left(0,\left(\sigma_{R}\right)^{2}\right)$, which is estimated for year $y 1$ to $y 2$ (see equation (B4)), and
$\sigma_{R} \quad$ is the standard deviation of the log-residuals, which is input ( $\sigma_{R}=0.4$ or $\sigma_{R}=1.5$ )

## B.2.5. Catches

$-\ell n L^{\text {Cach }}=\sum_{u} \sum_{y}\left[\frac{\ell n C_{u, y}-\ell n \hat{C}_{u, y}}{2 \sigma_{\mathrm{C}}^{2}}\right]$
where
$C_{u, y}$ is the observed catch of both species in year $y$ and Unit $u$,
$\hat{C}_{u, y}=\sum_{s} \hat{C}_{s, u, y}$ is the predicted catch in year $y$, and
$\sigma_{\mathrm{C}}$ is the input $\mathrm{CV}(=0.2)$.
No assumption about the species split of the catches is made on input; rather flexibility is allowed in the model by estimating the annual S. fasciatus proportion in the catches directly, by means of the following penalty added to the negative log-likelihood:
where
$-\ell n L^{F P p e n}=\sum_{u} \sum_{y=1960}^{2009}\left[\frac{\left(p_{y}^{u}-\mu^{u}\right)^{2}}{2\left(\sigma^{u}\right)^{2}}\right]$
$p_{y}^{u}$ is the estimated proportion of $S$. fasciatus in the catch in year $y$ and Unit $u$,
$\mu^{u}$ and $\sigma^{u}$ are the mean and standard deviation respectively of the distribution of $S$. fasciatus proportions in Unit $u$ based on the survey species split information (McAllister and Duplisea, 2012). For Unit 1, $\mu^{u}=0.40$ and $\sigma^{u}=0.16$, and for Unit $2, \mu^{u}=0.53$ and $\sigma^{u}=0.10$.

## B.3. Model parameters

## B.4.1. Fishing selectivity-at-length:

The survey fishing selectivity-at-length, $S_{s, u, l}^{s u r v}$ are estimated directly for a series of lengths (from 10 cm to 40 cm by 5 cm steps) and is taken to be linear between these lengths. The slope from lengths $I_{\text {minus }}$ to $I_{\text {minus }}+1$ is assumed to continue exponentially to lower lengths down to length 1. For lengths above $I_{\text {plus, }}$, the selectivity is taken to be flat (i.e. $S_{l}=S_{l_{\text {plus }}}$ for $1>I_{\text {plus }}$ ).

The commercial fishing selectivities-at-length, $S_{s, u, l}$ are estimated in terms of a logistic curve given by:

$$
\begin{equation*}
S_{s, u, l}=\frac{1}{1+\exp \left(-\left(l-l_{s, u}^{c}\right) / \boldsymbol{\delta}_{s, u}^{c}\right)} \tag{B28}
\end{equation*}
$$

where
$l_{s, u}^{c} \mathrm{~cm}$ is the length-at-50\% selectivity for species $s$ in Unit $\underline{u}$,
$\delta_{s, u}^{c} \mathrm{~cm}^{-1}$ defines the steepness of the ascending limb of the selectivity curve for species s in Unit $u$.

In practice however, the commercial selectivities have been taken to be the same for the two Units.

Table B1: Minus and plus length groups (in cm ) for the commercial and survey CAL.

|  | S. mentella | S. fasciatus |
| ---: | :---: | :---: |
| Commercial CAL: |  |  |
| $I_{\text {minus }}$ | 20 | 20 |
| $I_{\text {plus }}$ | 45 | 45 |
| Survey CAL: |  |  |
| $I_{\text {minus }}$ | 20 | 20 |
| $I_{\text {plus }}$ | 45 | 45 |

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